

	EXAMINATION QUESTIONS/SOLUTIONS 2007-08 E2.11 - ISE2 MATHS - SOLUTIONS 2008	Course <u>ISE2.</u>
Question <u>1</u>		Marks & seen/unseen
Parts	<p>(i) $\hat{u}_a(\omega) = \int_{-\infty}^{\infty} u_a(t) e^{-i\omega t} dt = \int_{-a}^a e^{-i\omega t} dt$ $= -\frac{1}{i\omega} [e^{-i\omega a} - e^{i\omega a}] = \underline{\underline{\frac{2}{\omega} \sin(\omega a)}}$</p> <p>(ii) Since $\frac{dh}{dt} = g$ we have $\hat{g} = i\omega \hat{h}$ & also $g(t) = f(t+a) - f(t-a)$ $\Rightarrow \hat{g}(\omega) = \int_{-\infty}^{\infty} f(t+a) e^{-i\omega t} dt - \int_{-\infty}^{\infty} f(t-a) e^{-i\omega t} dt$ $= \int_{-\infty}^{\infty} f(s) e^{-i\omega(s-a)} ds - \int_{-\infty}^{\infty} f(s) e^{-i\omega(s+a)} ds$ $= e^{i\omega a} \hat{f}(\omega) - e^{-i\omega a} \hat{f}(\omega)$ [or can just quote shift rule] $= 2i \sin(\omega a) \hat{f}(\omega)$ Thus we have $\hat{h} = \underline{\underline{\frac{2}{\omega} \sin(\omega a) \hat{f}}}$ as the req'd relation.</p> <p>(iii) Using convolution: $h(t) = (FT)^{-1} \left\{ \underbrace{\frac{2}{\omega} \sin(\omega a)}_{\hat{g}, \text{ say}} \hat{f} \right\}$ $= \int_{-\infty}^{\infty} f(t-s) g(s) ds$ but, from part (i) we see that $g(t) = u_a(t)$. $\therefore h(t) = \int_{-\infty}^{\infty} f(t-s) u_a(s) ds$ $= \underline{\underline{\int_{-a}^a f(t-s) ds}}$ [N.B. this is also equal to $\int_{-a}^a f(t+s) ds$]</p> <p>(iv) By differentiating this last expression we have $\frac{dh}{dt} = \int_{-a}^a f'(t-s) ds$ ← [or use $\int_{-a}^a f(t+s) ds$] $= \int_{t-a}^{t+a} f'(q) dq$ ($q=t-s$) $= f(t+a) - f(t-a)$ which is \hat{g}, so OK ✓</p>	<p>4</p> <p>6 mostly seen, but phrased differently</p> <p>6</p> <p>4 (Total 20)</p>
Setter's initials <u>D. Holm</u>	Checker's initials <u>Agw</u>	Page number

	EXAMINATION QUESTIONS/SOLUTIONS 2007-08	Course ISE 2.
Question 2		Marks & seen/unseen
Parts (i)	$\int_{-\infty}^{\infty} \hat{f}(\omega) ^2 d\omega = \int_{-\infty}^{\infty} \hat{f}(\omega) \hat{f}^*(\omega) d\omega$ $= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} f(t) e^{-i\omega t} dt \right) \left(\int_{-\infty}^{\infty} f^*(u) e^{i\omega u} du \right) d\omega$ <p>Changing the order of integration:</p> $= \int_{t=-\infty}^{\infty} f(t) \int_{u=-\infty}^{\infty} f^*(u) \left(\int_{\omega=-\infty}^{\infty} e^{-i\omega(t-u)} d\omega \right) du dt$ $= 2\pi \int_{t=-\infty}^{\infty} f(t) f^*(t) dt$ $= 2\pi \int_{-\infty}^{\infty} f(t) ^2 dt \quad \text{Hence result.}$ <p>Alternatively, this may be proved using the convolution theorem</p> <p>(ii) We are given $f(t) = \begin{cases} 0 & t < 0 \\ e^{-t} & t \geq 0 \end{cases}$</p> <p>LHS of Parseval formula is $\int_{-\infty}^{\infty} f(t) ^2 dt = \int_0^{\infty} e^{-2t} dt$</p> $= \left[-\frac{1}{2} e^{-2t} \right]_0^{\infty} = \frac{1}{2}$ <p>To evaluate RHS, first need $\hat{f}(\omega)$.</p> $\hat{f}(\omega) = \int_0^{\infty} e^{-t} e^{-i\omega t} dt = \int_0^{\infty} e^{-(1+i\omega)t} dt$ $= \left[-\frac{1}{1+i\omega} e^{-(1+i\omega)t} \right]_0^{\infty}$ $= \frac{1}{1+i\omega}$ <p>Then $\hat{f} ^2 = \hat{f} \hat{f}^* = \frac{1}{1+i\omega} \cdot \frac{1}{1-i\omega}$</p> $= \frac{1}{1+\omega^2}$ <p>Thus, RHS of Parseval is $\frac{1}{2\pi} \int_{-\infty}^{\infty} \frac{d\omega}{1+\omega^2}$ (formula sheet) $= \frac{1}{2\pi} [\tan^{-1}(\omega)]_{-\infty}^{\infty} = \frac{1}{2\pi} \left(\frac{\pi}{2} - \left(-\frac{\pi}{2}\right) \right) = \frac{1}{2}$</p>	<p>2</p> <p>2</p> <p>2</p> <p>2</p> <p>2</p> <p>2</p> <p>(i) is seen (ii) is unseen</p> <p>2</p> <p>1</p> <p>3</p> <p>2</p> <p>2</p> <p>(Total 20)</p>
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EXAMINATION QUESTIONS/SOLUTIONS 2007-08

Course

ISE 2

Question

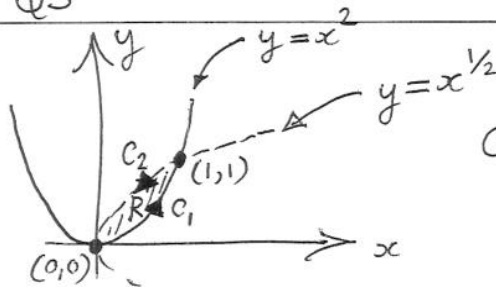
SOLUTION
To Q3

Marks &

seen/unseen

Parts

(i)



C is formed from the union of C_1 & C_2 .

C_1 goes from $(0,0)$ to $(1,1)$

C_2 goes from $(1,1)$ to $(0,0)$

5

(ii)

For $n=1$ we have $y=x^2$ & hence $dy=2x dx$.
Thus, writing I_1 in terms of x we have

$$I_1 = \int_0^1 \frac{2x^8}{2x^2} dx - \int_0^1 \frac{x^7 \cdot 2x dx}{2x^2} = \underline{\underline{0}}$$

3

For $n=2$ we have $y^2=x \Rightarrow 2y dy = dx$
Writing I_2 in terms of y we have

$$I_2 = \int_1^0 \frac{y^4}{(y^2)^2} \cdot 2y dy - \int_1^0 \frac{y^3}{y^2} dy$$

$$= \int_1^0 (2y - \frac{1}{2}y) dy = \frac{3}{2} \left[\frac{y^2}{2} \right]_1^0 = \underline{\underline{-\frac{3}{4}}}$$

4

(iii)

We have $P = y^4/x^2$, $Q = -\frac{1}{2}y^3/x$

$$\text{Thus } \frac{\partial Q}{\partial x} - \frac{\partial P}{\partial y} = \frac{1}{2} \frac{y^3}{x^2} - \frac{4y^3}{x^2} = -\frac{7}{2} \frac{y^3}{x^2}$$

\therefore Applying Green's Theorem:

$$\int_C P dx + Q dy = \iint_R -\frac{7}{2} \frac{y^3}{x^2} dx dy$$

$$= \int_{x=0}^1 \int_{y=x^2}^{y=x^{1/2}} -\frac{7}{2} \frac{y^3}{x^2} dy dx$$

$$= -\frac{7}{8} \int_{x=0}^1 \frac{1}{x^2} [y^4]_{x^2}^{x^{1/2}} dx$$

$$= -\frac{7}{8} \int_0^1 (1-x^6) dx = \underline{\underline{-\frac{3}{4}}}$$

unseen
but
similar
problems
done.

Total
20

6

(iv)

Since $\int_C \equiv \int_{C_1} + \int_{C_2}$, we have $-\frac{3}{4} = 0 - \frac{3}{4}$ \checkmark so results are consistent.

2

Setter's initials

Agw

Checker's initials

Page number

	EXAMINATION QUESTIONS/SOLUTIONS 2007-08	Course <u>ISE 2</u>
Question <u>4</u>		Marks & seen/unseen
Parts	<p>(i) Poles of $f(z)$ occur when denominator is zero. 1</p> <p>i.e. $z(3z^2+13z+4) = 0$</p> <p>$\Rightarrow z(3z+1)(z+4) = 0$</p> <p>$\Rightarrow z = 0, z = -\frac{1}{3}, z = -4$ 3</p> <p>To find order of poles:</p> <p>At $z=0$: Consider $\lim_{z \rightarrow 0} z^\alpha f(z) = \lim_{z \rightarrow 0} \frac{z^{\alpha-1}}{3z^2+13z+4}$</p> <p>Similarly: $\neq 0$ & finite only if $\alpha=1$.</p> <p>At $z=-\frac{1}{3}$: $\lim_{z \rightarrow -\frac{1}{3}} (z+\frac{1}{3})^\alpha f(z) \neq 0$ iff $\alpha=1$ i.e. order of pole is 1</p> <p>$z=-4$: $\lim_{z \rightarrow -4} (z+4)^\alpha f(z) \neq 0$ iff $\alpha=1$ So all poles are simple.</p> <p>[Alternatively just state that each zero of the denominator occurs to multiplicity one] 3</p> <p>Res $f(z) = \lim_{z \rightarrow 0} z f(z) = \frac{1}{4}$ 2</p> <p>Res $f(z) = \lim_{z \rightarrow -\frac{1}{3}} (z+\frac{1}{3}) \frac{1}{3z(z+\frac{1}{3})(z+4)} = \frac{-3}{11}$ 2</p> <p>Res $f(z) = \lim_{z \rightarrow -4} (z+4) \frac{1}{z(3z+1)(z+4)} = \frac{1}{44}$ 2</p> <p>(ii) Hence: $\mathcal{L}^{-1}(F(s)) =$ Sum of residues of $F(s)e^{st}$ at poles of $F(s)$ 1</p> <p>Res $F(s)e^{st} = \frac{1}{4}$; Res $F(s)e^{st} = -\frac{3}{11} e^{-t/3}$ $s=0$ $s=-\frac{1}{3}$</p> <p>Res $F(s)e^{st} = \frac{1}{44} e^{-4t}$ using results obtained in (i) $s=-4$</p> <p>Thus $\mathcal{L}^{-1}(F(s)) = \frac{1}{4} - \frac{3}{11} e^{-t/3} + \frac{1}{44} e^{-4t}$ 6</p> <p>[Otherwise: P.F.'s: $F(s) = \frac{1}{4s} + \frac{1/44}{s+4} - \frac{9/11}{(3s+1)}$ Invert, using tables: $f(t) = \frac{1}{4} + \frac{1}{44} e^{-4t} - \frac{3}{11} e^{-t/3}$]</p>	<p>Unseen but a similar problem studied in class</p> <p>2</p> <p>2</p> <p>2</p> <p>1</p> <p>6</p> <p>Total 20</p>
Setter's initials <u>AlgW</u>	Checker's initials <u>X.WU</u>	Page number

5. Let

T_1 - event first test passes (no memory errors detected)

\overline{T}_1 - event first test fails (memory errors detected)

T_2 - event second test passes (no memory errors detected)

\overline{T}_2 - event second test fails (memory errors detected)

E - event that there are memory errors

\overline{E} - event that there are no memory errors

(i)

$$\begin{aligned} P(T_1 | \overline{E}) &= 1 & P(\overline{T}_1 | \overline{E}) &= 0 \\ P(T_1 | E) &= 0.2 & P(\overline{T}_1 | E) &= 0.8 \\ P(T_2 | \overline{E}) &= 1 & P(\overline{T}_2 | \overline{E}) &= 0 \\ P(T_2 | E) &= 0.01 & P(\overline{T}_2 | E) &= 0.99 \\ P(E) &= 0.02 & P(\overline{E}) &= 0.98 \end{aligned}$$

(a)

$$\begin{aligned} P(\overline{T}_1) &= P(\overline{T}_1 | E)P(E) + P(\overline{T}_1 | \overline{E})P(\overline{E}) \\ &= 0.8 \times 0.02 + 0 \times 0.98 = 0.016. \end{aligned}$$

2

(b)

$$\begin{aligned} P(E | \overline{T}_1) &= \frac{P(\overline{T}_1 | E)P(E)}{P(\overline{T}_1)} \\ &= \frac{0.8 \times 0.02}{0.016} = 1. \end{aligned}$$

2

(c) (as expected!)

$$\begin{aligned} P(T_1 \cap T_2) &= P(T_1 \cap T_2 | E)P(E) + P(T_1 \cap T_2 | \overline{E})P(\overline{E}) \\ &= P(T_1 | E)P(T_2 | E)P(E) + P(T_1 | \overline{E})P(T_2 | \overline{E})P(\overline{E}) \\ &= 0.2 \times 0.01 \times 0.02 + 1 \times 1 \times 0.98 = 0.98004. \\ P(E | T_1 \cap T_2) &= \frac{P(T_1 \cap T_2 | E)P(E)}{P(T_1 \cap T_2)} \\ &= \frac{P(T_1 | E)P(T_2 | E)P(E)}{P(T_1 \cap T_2)} \\ &= \frac{0.2 \times 0.01 \times 0.02}{0.98004} = 4.08 \times 10^{-5}. \end{aligned}$$

6

Gm

YH

(5)

- (ii) Let R_1 = running time of first test
 R_2 = running time of second test

$$R_1 \sim N(5, 2^2) \quad R_2 \sim N(60, 10^2) \quad Z \sim N(0, 1)$$

(a)

$$\begin{aligned} P(R_1 > 6) &= P\left(\frac{R_1 - 5}{2} > \frac{6 - 5}{2}\right) = P\left(Z > \frac{1}{2}\right) \\ &= 1 - \Phi(0.5) = 1 - 0.691 = 0.309. \end{aligned}$$

3

(b)

$$\begin{aligned} P(R_1 < 3) &= P\left(\frac{R_1 - 5}{2} < \frac{3 - 5}{2}\right) = P(Z < -1) \\ &= \Phi(-1) = 1 - \Phi(1) = 1 - 0.841 = 0.159. \end{aligned}$$

2

(c) Total test time $R = R_1 + R_2$.

$$R \sim N(5 + 60, 2^2 + 10^2) = N(65, 104)$$

2

(d) Given

$$\bar{x} = 5.24, s = 2.12, t_0 = t_{n-1, 0.05} = 2.26.$$

95% CI for mean:

$$\begin{aligned} \left(\bar{x} - t_0 \frac{s}{\sqrt{n}}, \bar{x} + t_0 \frac{s}{\sqrt{n}}\right) &= \left(5.24 - 2.26 \times \frac{2.12}{\sqrt{10}}, 5.24 + 2.26 \times \frac{2.12}{\sqrt{10}}\right) \\ &= (3.725, 6.755) \end{aligned}$$

this interval contains the reported mean value of 5 minutes, so we would fail to reject a hypothesis that the mean is 5 minutes at the 5% level.

3

(20)

6m

4H

6. (i) (a)

$$\int_{-\infty}^{\infty} f(t) dt = \int_0^{\infty} \lambda \beta (\lambda t)^{\beta-1} e^{-(\lambda t)^\beta} dt = [-e^{-(\lambda t)^\beta}]_0^{\infty} = 1,$$

and $f(t) \geq 0 \forall t \geq 0$ as $\lambda, \beta \geq 0$.

3

(b) Reliability:

$$\begin{aligned} R(t) &= P(T > t) = \int_t^{\infty} \lambda \beta (\lambda t_0)^{\beta-1} e^{-(\lambda t_0)^\beta} dt_0 \\ &= [-e^{-(\lambda t_0)^\beta}]_t^{\infty} = e^{-(\lambda t)^\beta} \end{aligned}$$

2

Hazard:

$$h(t) = \frac{f(t)}{R(t)} = \frac{\lambda \beta (\lambda t)^{\beta-1} e^{-(\lambda t)^\beta}}{e^{-(\lambda t)^\beta}} = \lambda \beta (\lambda t)^{\beta-1}$$

2

(c)

$$\begin{aligned} P(T > t_0 + t | T > t_0) &= \frac{P(T > t_0 + t \cap T > t_0)}{P(T > t_0)} \\ &= \frac{P(T > t_0 + t)}{P(T > t_0)} = \frac{e^{-(\lambda(t+t_0))^\beta}}{e^{-(\lambda t_0)^\beta}} \\ &= e^{(\lambda t_0)^\beta - (\lambda(t+t_0))^\beta} \end{aligned}$$

3

(d) Now, $P(T > t) = e^{-(\lambda t)^\beta}$.

"memoryless" when $P(T > t_0 + t | T > t_0) = P(T > t)$. From part (c), this occurs when $\beta = 1$, (and, of course trivially when $t_0 = 0$!) giving $T \sim \text{Exponential}(\lambda)$.

2

(ii) Let R be the reliability at 30 minutes (=0.5 hours).

Let A, B_1, B_2 be the events that the corresponding component is operating at 30 minutes.

$$P(T > t) = e^{-(\lambda t)^\beta} \Rightarrow P\left(T > \frac{1}{2}\right) = e^{-\left(\frac{1}{2}\right)^\beta}$$

2

Giving,

$$P(A_1) = e^{-\left(\frac{0.5}{2}\right)^{0.8}} = 0.7190; \quad P(B_1) = P(B_2) = e^{-\left(\frac{0.5}{2}\right)^{0.5}} = 0.6065.$$

2

$$\begin{aligned} R &= P(A_1 \cap (B_1 \cup B_2)) = P(A_1)P(B_1 \cup B_2) \\ &= P(A_1)(P(B_1) + P(B_2) - P(B_1 \cap B_2)) \\ &= P(A_1)(P(B_1) + P(B_2) - P(B_1)P(B_2)) \\ &= 0.7190 \times (0.6065 + 0.6065 - 0.6065^2) = 0.6077. \end{aligned}$$

4

Gm

YH